Planning While Flying: A Measurement-Aided Dynamic Planning of Drone Small Cells

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Abstract—The deployment of drone small cells has emerged as a promising solution to agile provisioning of Internet backbone access for Internet of Things devices, and many other types of users/devices. In this paper, we consider the problem of deploying a set of drone cells operating on multiple channels in a target area to provide access to the backbone/core network, which is formulated as a combinatorial network utility maximization problem. Since an offline and centralized solution to such a problem is not feasible, a low-complexity and distributed online algorithm is highly desired. Therefore, we propose a measurement-aided dynamic planning (MAD-P) algorithm, where the dispatched drones perform position and channel configurations autonomously on the fly based on the real-time measurement of network throughput to solve the problem in a distributed fashion during flight with minimal centralized control. We prove that the proposed MAD-P algorithm is asymptotically optimal, and investigate how long it takes for the convergence to stationarity under the MAD-P algorithm by giving a mixing time analysis. We also derive an upper bound of the performance gap in presence of measurement errors. Simulation results are provided to validate our analytic results and demonstrate the effectiveness of our algorithm.

Index Terms—Distributed algorithm, drone small cell, Internet of Things (IoT), Markov approximation, optimal dynamic planning, unmanned aerial vehicles (UAVs).

I. INTRODUCTION

U NMANNED Aerial Vehicles (UAVs), commonly known as drones, have gained rapid development in both military and civilian domains. Recently, substantial interest has

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been attracted in the development of drone-based airborne communication networks for applications such as harvesting data from sensors deployed in hardly accessible areas [1]–[6], enhancing coverage of cellular networks [7], and aiding communications between the reader and radio frequency identifiers in battery-free networks [8], among others. Particularly, recent research [9] has demonstrated the feasibility of mounting a small cell base station (sBS) on a flying drone to extend the last-mile connectivity to ground users/Internet of Things (IoT) devices that require accessing the backbone/core network in a region of interest.

Unlike deployment of fixed access infrastructure which typically requires time-consuming network planning taking factors, such as propagation, geographic limitations, and traffic distribution (e.g., [10] and [11]) into considerations, and is difficult to change or to optimize over time, the main advantage of leveraging drone sBSs is that their deployment can be *agile* and *readily reconfigurable* due to the flexible mobility of drones. For example, a cluster of drone sBSs can be quickly launched regardless of geographical terrain; the position of drones can be adjusted in response to variations of wireless connectivity; and the cluster of operating drone sBSs can scale up and down in response to the change of network traffic demand. This has emerged as a promising solution to agile cellular/Internet service provisioning in many application scenarios (e.g., coverage and capacity enhancement of 5G/Beyond 5G cellular networks during temporary events [12], [13], offloading the computation-intensive tasks onto the cloud from mobile IoT devices [14], [15], etc.).

A. Drone Small Cells Planning Problem

In this paper, we consider the deployment of a set of drone sBSs in a target area of multiple cells and over multiple communication channels to serve the data traffic demand in the target area (i.e., a model of providing access to the backbone network), as shown in Fig. 1. The data traffic demand in each cell is assumed unknown (before deployment) and time-invariant but varies across cells.

1) Key Question to Answer: The question we would like to answer is that what is the best deployment (configuration) of drone sBSs such that the most data traffic demand can be served? This is not trivial because different configurations of drone sBSs in terms of cell placement and channel assignment may yield different system capacity and therefore the system throughput (i.e., the total demand being served). This is mainly

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Fig. 1. Network of drone sSBs serving a target area.

due to the mutual interference generated under a certain configuration. For example, intuitively, we intend to send three drones (all we have) to the three most demanding cells so that we could fully utilize the capacity. However, it happens that the three cells are very close to each other geographically and there is only one communication channel available (in an extreme case) such that the total serving capacity might be very limited and less that the total traffic demand of the three cells due to the strong mutual interference. This would lead to a suboptimal configuration, and there may exist other configurations that can provide higher system throughput.

2) Planning Problem: The goal of the planning is to dispatch drone sBSs to cells and assign channels to drones such that the maximum system throughput is achieved. It can be seen that the planning process is essentially a joint cell placement (trying to assign drones to serve the most demanding cells) and channel assignment (trying to distribute drones evenly over channels and cells to minimize interference so as to maximize the total serving capacity) process given the data traffic demand of all cells (unknown before deployment). We further generalize the system throughput (the total demand being served) maximization problem to a utility maximization problem where the utility is a concave and increasing function of the throughput, representing a certain utility of interest (e.g., service satisfaction).

3) Need of Low-Complexity Implementation: The drone sBSs planning problem is essentially a combinatorial optimization problem, i.e., finding an optimal configuration from the set of all possible configurations. However, since the data traffic demand is unknown before deployment (which is often the case in real-world scenarios), an offline and centralized solution (e.g., exhaustive search) is not feasible, and a low-complexity online solution is desired.

4) Need of Distributed Implementation: The desired planning algorithm should also be amenable to distributed implementations, such that computation is distributed over drones without having to be centralized at the backhaul gateway. Moreover, distributed implementations tend to be more scalable than centralized solutions. Also, due to the flexible mobility of drone sBSs, reconfiguration (moving to a different cell or switching to a different channel) can be agile and autonomous.

Therefore, we aim at solving the drone sBSs planning problem in a low-complexity and distributed way. In this paper, we propose an online measurement-aided dynamic planning (MAD-P) algorithm, where the dispatched drones perform self-configurations on the fly regarding the cell being located and the channel being used, based on the real-time measurement of network throughput. The proposed planning process is autonomous, and no prior knowledge of data traffic demand in the target area is required. The main contributions of this paper are highlighted as follows.

- The proposed algorithm leads to a joint planning of cell placement and channel selection, solves the underlining utility maximization problem by using the Markov approximation techniques, and is amenable to lowcomplexity and distributed implementation. In addition, we introduce a design parameter to control the preference of the exploration of cells over the exploration of channels.
- 2) We prove the asymptotic optimality of the proposed algorithm. To understand how long it takes for the convergence to stationarity under the MAD-P algorithm, we derive an upper bound of the mixing time that captures the speed of convergence of the resulting Markov chain. We also show that the performance gap in presence of measurement errors is upper bounded.

B. Literature Review

The deployment of drone sBSs also faces great challenges. Key implementation issues have been investigated by recent studies, such as: 1) due to the lack of fiber-based connectivity, a viable wireless backhauling/fronthauling solution should be in place. The state-of-the-art solutions resort to different radio technologies, including millimeter wave [16]–[18], satellite links [19], and free-space optics [20], each of which has its own pros and cons and 2) due to the short endurance of drones on a single charge, it is crucial to prolong the service duration or even provide persistent service during the mission via advanced charging technologies. The state-of-the-art solutions include equipping drones with solar panels [21], charging the battery during flight by using a high energy laser beam [22], and charging the battery by more powerful fixed-wing UAVs via wireless power transfer [23]. Moreover, some theoretical perspectives on efficient deployment of drone sBSs have also been provided in the literature.

- 1) In [24], the optimal altitude of deployment is derived in maximizing the ground coverage.
- 2) The service time maximization considering the traveling time of mobile small cells is given in [25].
- Dynamic repositioning of drone sBSs is considered in [26] to improve the spectral efficiency.
- 4) The optimal placement of drone BSs considering energy efficiency is given in [27].
- Algorithms for improving the communication throughput of a UAV network are proposed (e.g., [28] and [29]).
- Channel assignment, or generally radio resource allocation, is considered in [30] and [31].

However, despite these advances, a joint cell placement and channel assignment planning process toward system throughput maximization has not been considered. Moreover, a low-complexity and distributed online algorithm relying on dynamic and autonomous configurations of drone sBSs should be in place toward the optimal planning, which has not been reported in existing research works either.

The remainder of this paper is organized as follows. Section II introduces the system model. The problem formulation is given in Section III. Section IV presents the proposed dynamic drone cell planning algorithm, followed by the performance analysis in Section V. In Section VI, we provide the simulation results and Section VII concludes this paper.

II. SYSTEM MODEL

We consider deploying a set of drone sBSs, denoted by $\{1, 2, \ldots, N\} \triangleq [N]$, to a rectangular target area consisting of $L \times W$ square cells, each of which is r^2 in size. Each square cell, indexed by (l, w), where $l \in \{1, 2, \dots, L\} \triangleq [L]$ and $w \in \{1, 2, \dots, W\} \triangleq [W]$, can be fully covered by a drone sBS at a given height H. Let \mathcal{Q} denote the set of $L \times W$ distinct cells, i.e., $Q = \{(l, w) : l \in [L], w \in [W]\}$, and $q_n = (l_n, w_n)$ denote the cell where drone sBS n is deployed, where $l_n \in [L]$ and $w_n \in [W]$. We consider that a cell is at most served by one drone sBS, and hence multiple drone sBSs are not allowed to be coexisting in one cell. All drone sBSs are assumed directly connected to a central station via point-to-multipoint microwave/mmWave backhauling with sufficient bandwidth. The central station serves as a gateway to the backbone network and a hub for exchanging messages among drone sBSs, as shown in Fig. 1. For simplicity, we consider that each drone can be replaced in time to maintain system operation. A drone

will monitor its battery status periodically. If the battery level is lower than a preset threshold, the drone will send a message to the control station to schedule a fully charged drone for replacement.

A multichannel system is considered to reduce the cochannel interference among ground-air links in different cells. we consider that each drone sBS can select a channel to operate on from a set of M nonoverlapping channels of equal bandwidth, denoted by $\{1, 2, ..., M\} \triangleq [M]$. The channel selected by drone sBS n can be represented as c_n , where $c_n \in [M]$. Moreover, we consider that the number of channels is limited and no more than the number of dispatched drones, i.e., M < N. Since each drone sBS is functioning in a certain cell under a certain channel at any given time, we represent by the drone cell planning process $\{\psi(t)\}_{t>0}$. At any time $t, \psi(t)$ takes a value f from a finite set \mathcal{F} , where $f \triangleq ((q_n, c_n))_{n \in [N]}$ represents a certain planning/configuration of all drone sBSs in the system and \mathcal{F} denotes the set of all possible configurations.

We assume that a number of long-lived data flows are distributed in the target area unevenly, leading to heterogeneous data traffic demand in cells. Each long-lived data flow is a traffic stream that is always in the network and continually generates bits at a certain rate. The data traffic demand in the target area is denoted by $\mathbf{D} = (d_{l,w})_{l \in [L],w \in [W]}$, where $d_{l,w}$ is the aggregate rate generated from all data traffic flows in cell (l, w). In this paper, the data traffic demand matrix \mathbf{D} is assumed time-invariant. However, as what we will discuss in detail in Section IV-D, the optimality of the proposed algorithm remains if \mathbf{D} changes slowly with time.

Given the drone cell planning $\psi(t)$ at time *t*, a certain capacity vector $\mathbf{S}(t) = (s_n^{\psi(t)})_{n \in [N]}$ is obtained, where $s_n^{\psi(t)}$ is the maximum data rate at which the data traffic demand is served by drone sBS n, and is determined by the interference pattern under $\psi(t)$ assuming a uniform ground-air transmission power over all cells. Therefore, the actual data demand served by (i.e., the ground-air throughput of) drone sBS *n* under $\psi(t)$ is given by $\gamma_n^{\psi(l)} = \min\{s_n^{\psi(l)}, d_{l_n, w_n}\}$, i.e., the throughput of a drone cell is the minimum of the traffic demand and service capacity, and we denote the ground-air throughput vector by $\Gamma(t) = (\gamma_n^{\psi(t)})_{n \in [N]}$. Note that the unserved traffic demand will be either ignored or delivered by other network infrastructures (e.g., cellular macro cells and satellite) if available. Also note that $\gamma_n^{\psi(t)}$ can be measured by drone sBS *n* in real time. That is the reason why the prior knowledge of **D** is not required for our algorithm design. We will discuss the affect of measurement errors on the proposed algorithm in Section V.

III. PROBLEM FORMULATION

The drone cell planning process $\{\psi(t)\}_{t>0}$ is essentially a joint *cell placement* and *channel assignment* process for all drone sBSs over time. The object of the planning process is to choose a configuration *f* from \mathcal{F} given **D** such that the maximum system throughput is achieved, i.e., maximizing $\|\Gamma(t)\|_1$, where $\|\cdot\|_1$ stands for the l_1 -norm. However, we generalize the problem to a utility maximization problem in terms of $\Gamma(t)$. Specifically, we consider the problem of finding the optimal

DDP:
$$\max_{f \in \mathcal{F}} \quad \frac{1}{N} \sum_{n \in [N]} \Theta_n \left(\gamma_n^f \right)$$

where $\Theta_n(\gamma_n^f)$ is the utility obtained with respect to throughput γ_n^f by drone sBS *n* under configuration *f*. The utility functions Θ_n for all *n* are assumed to be twice differentiable, strictly increasing and concave, and bounded by finite constants Θ_{\min} and Θ_{\max} , i.e.,

$$\Theta_{\min} \leq \Theta_n \left(\gamma_n^f \right) \leq \Theta_{\max} \quad \forall \ n \in [N] \text{ and } f \in \mathcal{F}.$$

As it can be seen, a planning policy (static) that selecting f^* by solving the combinatorial problem **DDP** is optimal. However, such a policy tends to be computationally prohibitive since the size of \mathcal{F} grows exponentially fast as the number of dispatched drones increases. The computational complexity of such a planning policy also depends on the number of available channels and the size of the target region. Therefore, our goal is to design a planning policy such that: 1) the computational complexity is significantly reduced and 2) drone sBSs make planning decisions distributedly without the centralized coordination.

IV. DYNAMIC DRONE CELL PLANNING

In this section, we first propose a Markov approximation to the original combinatorial optimization problem such that a static time-sharing policy is obtained. Then, we design a dynamic and distributed algorithm to implement the timesharing policy by constructing a continuous time-reversible Markov chain over all possible configurations.

A. Markov Approximation

Considering that the original problem **DDP** is hard to solve, inspired by [32], we obtain the following optimization problem by applying the Markov approximation techniques to the original problem:

$$\mathbf{DDP} - \mathbf{MA}: \quad \max_{p \ge 0} \quad \frac{p_f}{N} \sum_{n \in [N]} \Theta_n \left(\gamma_n^f \right) - \frac{1}{\beta} \sum_{f \in \mathcal{F}} p_f \log p_f$$

subject to
$$\sum_{f \in \mathcal{F}} p_f = 1$$

where $\boldsymbol{p} = (p_f)_{f \in \mathcal{F}}$ and p_f represents the fraction of time that configuration *f* is being used in the process $\{\psi(t)\}_{t>0}$; β is a positive constant and will be discussed later. The optimal solution \boldsymbol{p}^* to this problem is given by

$$p_{f}^{*} = \frac{\exp\left(\frac{\beta}{N}\sum_{n\in[N]}\Theta_{n}\left(\gamma_{n}^{f}\right)\right)}{\sum_{f'\in\mathcal{F}}\exp\left(\frac{\beta}{N}\sum_{n\in[N]}\Theta_{n}\left(\gamma_{n}^{f'}\right)\right)} \quad \forall f\in\mathcal{F}.$$
 (1)

This corresponds to a time-sharing policy where each configuration f is selected according to its time fraction p_f^* . Intuitively, a configuration f that leads to the maximum system-wide normalized utility is being used most often as it has the largest time fraction among all configurations, and the performance gap between such a time-sharing policy and the optimal static policy would be closed as β tends to infinity. It is worth noting that solving both **DDP** and **DDP** – **MA** problems require the knowledge of the throughput vector $\Gamma(f) = (\gamma_n^f)_{n \in [N]}$ for each f, which depends on the accurate modeling of the capacity vector $\mathbf{S}(f) = (s_n^f)_{n \in [N]}$ under a certain configuration f and a prior knowledge of **D**. This could be very difficult and computationally infeasible especially when **D** becomes time-variant. Considering that the throughput vector under a certain configuration can be obtained by the real-time measurement, we design an online measurementbased algorithm to implement the time-sharing policy given by (1) in a distributed manner.

B. Distributed Algorithm Design in General

The idea of distributed algorithm design is that we first construct a continuous time reversible Markov chain with the state space being \mathcal{F} and the desired stationary distribution given by (1); second, we design a distributed algorithm to realize the state transitions that drive the Markov Chain. Next, we will present the proposed MAD-P algorithm.

In the proposed algorithm design, a drone sBS may change its configuration by moving to a new cell or hopping to a new channel. Note that we do not consider the traveling time of drones which is negligible compared to the time scale where reconfigurations occur (e.g., at a time scale of minutes).¹ Further, we assume the channel switching time (e.g., a few milliseconds) is negligible as well. Note that there is energy cost (due to movements) and slightly performance reduction (due to service discontinuity during cell reconfiguration) of repositioning drones. That could be carefully modeled and quantified. However, considering the cost of repositioning drones brings an additional layer of complexity in the algorithm design. It would be more moderate to consider it in our future works.

Since changing either the cell placement or the channel assignment of a single drone sBS out of *N* is sufficient to lead to a new configuration, with the MAD-P algorithm employed, a configuration *f* of drone sBSs may change to another configuration *f'* at time *t* by having only one of drone sBSs either moving to a new cell or hopping to a new channel. Given an $f \in \mathcal{F}$, the cell placement and channel assignment of drone sBS *n* can be represented as q_n^f and c_n^f , respectively, where $q_n^f = (l_n^f, w_n^f), \ l_n^f \in [L], \ w_n^f \in [W], \ and \ c_n^f \in [M]$. It is possible to change configuration *f* to some other configuration *f'* if there exists an *n*₁, such that regarding the following conditions:

$$c1: q_{n_1}^{f} = q_{n_1}^{f'} \text{ and } c_{n_1}^{f} \neq c_{n_1}^{f'}; c2: q_{n_1}^{f} \neq q_{n_1}^{f'} \text{ and } c_{n_1}^{f} = c_{n_1}^{f'}; c3: \forall n_2 \in [N] \setminus n_1, q_{n_1}^{f'} \neq q_{n_2}^{f};$$

 $c4: \forall n_2 \in [N] \setminus n_1, q_{n_2}^f = q_{n_2}^{f'}$ and $c_{n_2}^f = c_{n_2}^{f'}$ either c1 and c4 hold (corresponding to a single drone sBS hopping to a new channel) or c2, c3, and c4 hold (corresponding to a single drone sBS moving to a new cell that is not being occupied/served).

¹We consider later to limit the mobility of drones for minimizing the impact of traveling time.

We denote by Q_n^f the set of cells that drone sBS *n* could possibly move to under current configuration *f*, which is dependent only on *f* and *n*. In the general design, we consider an *arbitrary mobility pattern*, where Q_n^f for each *f* and each *n* is an arbitrary set of unoccupied cells under *f* as long as the following two constraints hold:

c5: for each f and each n, $\forall q' \in Q_n^t$, if f' is due to that n moves from cell q to q', then $q \in Q_n^{f'}$;

c6: $Q(N) \subset \{\{q_n^I\}_{n \in [N]}\}_{f \in \mathcal{F}}$, where Q(N) is the set containing all *N*-combination sets² of Q. c5 states that it is always possible for each drone to move directly back to the old cell after moving to the new cell; while c6 states that the set of cell configurations of *N* drones due to some arbitrary mobility pattern should include all possible cell configurations, each of which is a set of *N* distinct cells chosen from Q. Note that $\{\{q_n^f\}_{n \in [N]}\}_{f \in \mathcal{F}}$ may contain permutations of one or more *N*-combination sets of Q. We consider that all permutations of the same *N*-combination set of *Q* with exactly the same channel configuration of drones lead to the same system utility, given that all drones are identical.

Moreover, let C_n^f be the set of channels that are not used by drone sBS *n* under current configuration *f*, i.e., $C_n^f = [M] \setminus c_n^f$. With the MAD-P algorithm employed, at a certain time instant, one of drone sBSs, say *n*, under configuration *f* would either move to a new cell in Q_n^f or hop to a new channel in C_n^f probabilistically, leading to a transition from *f* to a new configuration, say *f'*. The MAD-P algorithm is described in Algorithm 1.

In the proposed algorithm, we introduce a fixed parameter $\zeta \in (0, 1)$ to control the transitions between configurations. Clearly, increasing ζ increases the chance of seeing a cell reconfiguration; while decreasing ζ increases the chance of seeing a channel hopping. Under different system settings, we may want to choose different values of ζ . For example, in the case where we have $M \approx N \ll L \times W$, a value of ζ approaching one is preferable to explore the target area as much as possible to identify the most demanding cells to serve. In the extreme case where $\zeta = 1$, our algorithm reduces to a pure cell placement algorithm; while in the extreme case where $\zeta = 0$, our algorithm reduces to a pure channel assignment algorithm. Therefore, the MAD-P algorithm is considered as a randomized algorithm unifying both cell placement and channel assignment processes.

Proposition 1: With the MAD-P algorithm employed, the drone cell planning process $\{\psi(t)\}_{t>0}$ is a continuous-time time-reversible ergodic Markov chain with the stationary distribution given by (1).

Proof: It can be seen that $\{\psi(t)\}_{t>0}$ is a continuous-time homogeneous Markov chain since the amount of time the planning process stays in configuration/state *f* is exponentially distributed and dependent only on state *f*, and the transition probabilities are independent of time. Moreover, $\{\psi(t)\}_{t>0}$ is irreducible since it is always possible to transit from state *f* to

Algorithm 1 MAD-P Algorithm

- **Input:** The set of drone sBSs [N], the set of cells $((l, w))_{l \in [L], w \in [W]}$, the set of channels [M], the utility functions Θ_n for all $n \in [N]$, and predefined parameters $\zeta \in (0, 1), \alpha > 0, \beta > 0$, and H.
- **Output:** Drone cell planning process $\{\psi(t)\}_{t>t_0}$.

1: initialization:

- 2: At $t = t_0$, arbitrarily dispatch N drones to N distinct cells at the given height H and each drone randomly and uniformly selects a channel to operate on. This leads to an initialized configuration $f_0 \in \mathcal{F}$. Each drone sBS n obtains its ground-air throughput $\gamma_n^{f_0}$ via measurement and informs the central station of $(q_n^{f_0}, c_n^{f_0})$ and $\Theta_n(\gamma_n^{f_0})$. Then, the central station broadcasts f_0 and $(\Theta_n(\gamma_n^{f_0}))_{n \in [N]}$ to all drone sBSs. We have $\{\psi(t) = f_0\}_{t_0 < t \le t_1}$, where t_1 is the time when the 1st reconfiguration occurs.
- 3: while i > 0 (i^{th} reconfiguration) do
- 4: for $n \in [N]$ do

5:

As soon as drone sBS *n* receives f_{i-1} and $(\Theta_n(\gamma_n^{f_{i-1}}))_{n \in [N]}$ from the central station, it finds $Q_n^{f_{i-1}}$ and $C_n^{f_{i-1}}$, and counts down according to a generated random number from an exponential distribution with mean equal to

$$\frac{\exp\left(\frac{\beta}{N}\sum_{n\in[N]}\Theta_n(\gamma_n^{f_{i-1}})\right)}{\alpha\left(\zeta|\mathcal{Q}_n^{f_{i-1}}|+(1-\zeta)|\mathcal{C}_n^{f_{i-1}}|\right)};$$
(2)

6: end for

- 7: At $t = t_i$, the countdown of drone sBS n^* expires first among all drone sBSs. Drone sBS n^* immediately informs other drones via the central station to terminate their countdown processes, and does the following:
- 8: With probability $\frac{\zeta |Q_{n^*}^{f_{i-1}}|}{\zeta |Q_{n^*}^{f_{i-1}}| + (1-\zeta)|C_{n^*}^{f_{i-1}}|}$, drone n^* randomly and uniformly moves to one of the cells in $Q_{n^*}^{f_{i-1}}$; while with probability $\frac{(1-\zeta)|C_{n^*}^{f_{i-1}}|}{\zeta |Q_{n^*}^{f_{i-1}}| + (1-\zeta)|C_{n^*}^{f_{i-1}}|}$, drone n^* randomly and uniformly switches to one of the channels in $C_{n^*}^{f_{i-1}}$. 9: for $n \in [N]$ do
- 10: With the new configuration f_i , drone sBS *n* obtains its ground-air throughput $\gamma_n^{f_i}$ via measurement and informs the central station of $(q_n^{f_i}, c_n^{f_i})$ and $\Theta_n(\gamma_n^{f_i})$.
- 11: end for
- 12: The central station immediately broadcasts f_i and $(\Theta_n(\gamma_n^{f_i}))_{n \in [N]}$ to all drone sBSs.
- 13: Let $i \leftarrow i + 1$, and we have $\{\psi(t) = f_{i-1}\}_{t_{i-1} < t \le t_i}$. 14: **end while**

any other state f' in some finite time (either directly or indirectly with \mathcal{F} being finite and due to c5). Therefore, $\{\psi(t)\}_{t>0}$ has a unique stationary distribution from [33]. Next, we will show that $\{\psi(t)\}_{t>0}$ is time-reversible such that according to [34] the stationary distribution of $\{\psi(t)\}_{t>0}$ is given by (1).

To do so, it suffices to show that for any $f \in \mathcal{F}$ and any f' that f could possibly transit to, under Algorithm 1, the detailed

²A *k*-combination of a set S is a subset of *k* distinct elements of S.

balance equation holds, i.e.,

$$p_f^* \lambda (f \to f') = p_{f'}^* \lambda (f' \to f)$$
(3)

where $\lambda(f \to f')$ and $\lambda(f' \to f)$ are positive transition rates between state *f* and *f'*.

In Algorithm 1, since each drone sBS *n* counts down independently with a rate $[(\alpha(\zeta | Q_n^f| + (1 - \zeta) | C_n^f|))/(\exp((\beta/N) \sum_{n \in [N]} \Theta_n(\gamma_n^f)))]$, denoted by λ_n^f , in state *f*, the drone cell planning process will leave state *f* upon one of countdown processes expires at a rate of $\sum_{n \in [N]} \lambda_n^f$. Without loss of generality, we assume that the countdown process of drone sBS *n*^{*} expires first among all drone sBSs, i.e., the transition from *f* to *f'* is due to that drone sBS *n*^{*} either moves to a new cell or hops to a new channel. Since the length of each countdown process follows an individual exponential distribution, the probability that the countdown of drone sBS *n*^{*} expires first is given by $[(\lambda_{n^*}^f)/(\sum_{n \in [N]} \lambda_n^f)]$. Considering the case where *f* and *f'* differ only on the cell configuration of drone sBS *n*^{*} moves to $q_{n^*}^{f'}$ in $Q_{n^*}^f$ with probability $[(\zeta | Q_{n^*}^f|)/(\zeta | Q_{n^*}^f| + (1 - \zeta) | C_{n^*}^f|)] \cdot [1/(| Q_{n^*}^f|)]$, we have

$$p_{f}^{*}\lambda(f \to f') = \frac{\exp\left(\frac{\beta}{N}\sum_{n \in [N]}\Theta_{n}\left(\gamma_{n}^{f}\right)\right)}{\sum_{f'' \in \mathcal{F}}\exp\left(\frac{\beta}{N}\sum_{n \in [N]}\Theta_{n}\left(\gamma_{n}^{f''}\right)\right)}\left(\sum_{n \in [N]}\lambda_{n}^{f}\right)} \times \left(\frac{\lambda_{n^{*}}^{f}}{\sum_{n \in [N]}\lambda_{n}^{f}}\right)\frac{\zeta}{\zeta|\mathcal{Q}_{n^{*}}^{f}| + (1-\zeta)|\mathcal{C}_{n^{*}}^{f}|} = \frac{\alpha\zeta}{\sum_{f'' \in \mathcal{F}}\exp\left(\frac{\beta}{N}\sum_{n \in [N]}\Theta_{n}\left(\gamma_{n}^{f''}\right)\right)}.$$
(4)

Since state f' transits to state f by changing $q_{n^*}^{f'}$ to $q_{n^*}^{f}$, similarly, we have

$$p_{f'}^* \lambda \left(f' \to f \right) = \frac{\alpha \zeta}{\sum_{f'' \in \mathcal{F}} \exp\left(\frac{\beta}{N} \sum_{n \in [N]} \Theta_n\left(\gamma_n^{f''}\right)\right)}.$$
 (5)

By (4) and (5), we have (3). Now consider the case where f transits to f' by changing $c_{n^*}^f$ to $c_{n^*}^{f'}$, and f' transits to f by changing $c_{n^*}^{f'}$ to $c_{n^*}^{f}$. Similarly, we can prove that

$$p_{f}^{*}\lambda(f \to f') = p_{f'}^{*}\lambda(f' \to f)$$
$$= \frac{\alpha(1-\zeta)}{\sum_{f'' \in \mathcal{F}} \exp\left(\frac{\beta}{N} \sum_{n \in [N]} \Theta_{n}\left(\gamma_{n}^{f''}\right)\right)}.$$
 (6)

Proposition 1 shows that with the MAD-P algorithm, we achieve the desired stationary distribution over configurations such that the time-sharing policy according to (1) is implemented in a distributed manner where each drone sBS makes the decision locally on activating either a channel reconfiguration or a cell reconfiguration.

C. Two Special Designs

Next, we propose two special designs based on the MAD-P algorithm by specifying the mobility pattern of drones whenever they reconfigure their cell locations. 1) Maximum Mobility: In this special design, at any time a drone needs to change its cell placement, the drone can reach any unserved cell in the target area. Specifically, we consider that for each configuration f and each drone sBS n, Q_n^f contains all unoccupied cells, i.e., $Q_n^f = \{q : q \in Q, q \neq q_{n'}^f, \forall n' \in [N]\}$, which leads to the MAD-P algorithm with Maximum Mobility, called the MAD-P/MM algorithm. Note that actually given a configuration f, all drone sBSs have the same Q_n^f , denoted by Q^f , and we have $|Q^f| = LW - N$.

2) Limited Mobility: In this special design, at any time a drone needs to change its cell placement, the drone can only move to one of the unoccupied adjacent cells. Specifically, we consider that for each configuration f and each drone sBS n, $Q_n^f = \{q = (l, w) : |l - l_n^f| + |w - w_n^f| = 1, q \in Q, q \neq q_{n'}^f, \forall n' \in [N]\}$, such that each drone randomly and uniformly walks to one of the unoccupied adjacent cells (at most four) with a traveling distance of r. This gives us the MAD-P algorithm with limited mobility, called the MAD-P/LM algorithm. Limiting the mobility of drones is one way of minimizing the impact of traveling time since a drone sBS only needs to move to one of the neighboring cells (unserved) if the cell configuration needs to be changed.

D. Discussion on Implementations

1) Exploration Versus Exploitation: In the MAD-P algorithm, the time spent in a certain configuration f is exponentially distributed with a rate equal to

$$\sum_{n \in [N]} \lambda_n^f = \frac{\zeta \sum_{n \in [N]} |\mathcal{Q}_n^f| + (1 - \zeta) \sum_{n \in [N]} |\mathcal{C}_n^f|}{\frac{1}{\alpha} \exp\left(\frac{\beta}{N} \sum_{n \in [N]} \Theta_n\left(\gamma_n^f\right)\right)}.$$
 (7)

As it can be seen, the configuration that leads to a larger system-wide normalized utility tends to stay longer before it transits to other configurations. We can even prolong the sojourn time of a configuration (on average) by decreasing α such that we could potentially exploit a good configuration as much as possible when it occurs. We can also increase α to speed up the transitions among configurations. This particularly works for the scenario where the data traffic demand varies a lot across different cells since the algorithm should be able to identify the most demanding cells very quickly (i.e., the more exploration, the better). It is worth noting that tuning α does not change the stationary distribution of the planning process, and only affects the transition rates. As discussed before, the stationary distribution can be adjusted by changing β from (1). For example, we can simply increase the value of β to increase the time fraction that the system spent in the optimal configuration f^* . However, this could slow down the exploration over configurations according to (7). In a nutshell, the parameters α and β are used to tune the short-run tradeoff between exploration and exploitation of configurations in the planning process.

2) Message Passing: Note that the proposed algorithm depends on message passing among drone sBSs with the central station involved. It is associated with state transitions and is necessary to facilitate the decision making process at each drone sBS. The overhead of message passing will be reduced

in terms of *messages per unit time* if we slow down the state transitions in the planning process by tuning the parameters. In real-world implementations, the overhead introduced by message passing can be further reduced via the proper design of communication protocols.

3) *Time-Varying Traffic Demand:* The data traffic demand matrix **D** is assumed time-invariant. However, the optimality of the proposed algorithm can remain if **D** changes slowly with time. In addition, we need to choose an increasing function $g(\cdot)$ and reformulate the **DDP** problem as $\max_{f \in \mathcal{F}} g((1/N) \sum_{n \in [N]} \Theta_n(\gamma_n^f))$. For example, $g(\cdot) = \log \log(\cdot)$. Allowing **D** changes slowly actually corresponds to the relaxation of time-scale separation assumption in the research of queue-length-based CSMA algorithm, which has been well studied in [35].

V. PERFORMANCE ANALYSIS

In this section, we first prove that the proposed MAD-P algorithm is asymptotically optimal to the original **DDP** problem. Then, we will consider the problem of bounding the time taken by the planning process $\{\psi(t)\}_{t>0}$ to reach the desired stationary distribution under the MAD-P/MM algorithm. We leave the mixing time analysis for the MAD-P/LM algorithm and the general MAD-P algorithm for future works. Finally, we will discuss the impact of measurement errors on the algorithm performance.

A. Asymptotic Optimality

Theorem 1: The MAD-P algorithm is asymptotically utility-optimal to the **DDP** problem, i.e., for any $\epsilon \in (0, 1)$, $\delta \in (0, 1)$, there exists a constant C > 0, such that whenever $U(f^*) > C$, we have

$$\mathbf{Pr}\left\{U(f) > (1-\epsilon)U(f^*)\right\} > 1-\delta \tag{8}$$

where $U(f) = (\beta/N) \sum_{n \in [N]} \Theta_n(\gamma_n^f)$ is the weighted systemwide normalized utility under configuration f at time t and $U(f^*) = \max_{f \in \mathcal{F}} U(f)$.

Proof: Let \mathcal{I} denote the set of configurations that generate a weighted system-wide normalized utility less than $(1 - \epsilon)U(f^*)$, i.e.,

$$\mathcal{I} = \left\{ f \in \mathcal{F} : U(f) < (1 - \epsilon) U(f^*) \right\}.$$

Recall that p_f^* is the probability that configuration f is being used in steady state under the proposed MAD-P algorithm. Let $Z = \sum_{f \in \mathcal{F}} \exp(U(f))$. Then, we have

$$\Pr\{U(f) < (1 - \epsilon)U(f^*)\}$$

$$= \sum_{f \in \mathcal{I}} p_f^*$$

$$= \sum_{f \in \mathcal{I}} \frac{1}{Z} \exp(U(f))$$

$$\leq \frac{|\mathcal{I}|}{Z} \exp((1 - \epsilon)U(f^*))$$

$$< |\mathcal{I}| \exp(-\epsilon U(f^*))$$

$$< \frac{(LW)!}{(LW - N)!} M^N \exp(-\epsilon U(f^*))$$

where the last two inequities are due to the fact that $\exp(U(f^*)) < Z$ and $\mathcal{I} < |\mathcal{F}| \leq ((LW)!/(LW - N)!)M^N$,

respectively. As it can be seen, if we let $C = (1/\epsilon)(N \ln M + \ln ((LW)!/(LW - N)!) + \ln (1/\delta))$, and whenever $U(f^*) > C$, we have (8). Hence, the theorem holds.

Note that in theory Theorem 1 also provides a guideline for the design of utility functions $(\Theta_n)_{n \in [N]}$ for drone sBSs. Once parameters ϵ and δ are fixed, the constant C can be figured out such that $(\Theta_n)_{n \in [N]}$ can be designed wisely to ensure $U(f^*) > C$ under the optimal configuration f^* . In practice, it is difficult to obtain $U(f^*)$ before we actually run the algorithm since the ground-air throughput vector $\Gamma(f) = (\gamma_n^f)_{n \in [N]}$ needs to be measured on the go. Furthermore, we do not have the knowledge that which configuration is optimal until all the configurations in \mathcal{F} have been explored. One practical strategy is that we can run the algorithm for a short period of time T, find out $f_{\max} = \arg \max_{\psi(t), t \in (0,T]} U(\psi(t))$, adjust $(\Theta_n)_{n \in [N]}$ to have $U(f_{\text{max}}) > C$, and then rerun the algorithm. As a result of doing this, we must have $U(f^*) > C$. On the other hand, since $U(f) \geq \beta \Theta_{\min}, \forall f \in \mathcal{F}, \text{ we could simply have } \Theta_{\min} > C/\beta$ when we design $(\Theta_n)_{n \in [N]}$.

B. Speed of Convergence to Stationarity

The time sharing among configurations according to (1) is achieved once the resulting Markov chain enters its steady state (or equivalently, reaches its stationary distribution p^*). Therefore, it is important to understand how long it takes for the convergence to stationarity under the MAD-P/MM algorithm, and how different parameters affect the convergence behavior. Thus, we provide an upper bound of the mixing time that captures the speed of convergence of the resulting Markov chain $\{\psi(t)\}_{t>0}$ under the MAD-P/MM algorithm.

The mixing time (convergence time) of $\{\psi(t)\}_{t>0}$ is defined as follows:

$$\tau_{\min}(\epsilon) \triangleq \inf \left\{ t \ge 0 : \max_{f \in \mathcal{F}} d_{\mathrm{TV}} \big(\mathbf{P}_{f, \cdot}(t), \boldsymbol{p}^* \big) \le \epsilon \right\}$$

where $d_{\text{TV}}(\cdot, \cdot)$ denotes the total variation distance of two probability distributions, and $\mathbf{P}_{f,\cdot}(t)$ is the probability distribution of $\psi(t)$ over \mathcal{F} at time t given that the initial state is f.

1) Uniformization: To facilitate the mixing time analysis, motivated by [32], we follow the technique of *uniformization* in [36] to obtain a uniformization version of the continuoustime Markov chain $\{\psi(t)\}_{t>0}$, which is characterized by an embedded discrete-time Markov chain Ψ and a Poisson process with rate ρ . By [36, Th. 3.4], the original Markov chain and its uniformization version have the same transition rates between states.

We denote by $\mathbf{Q} = (Q_{ff'})_{|\mathcal{F}| \times |\mathcal{F}|}^3$ the transition rate matrix of $\{\psi(t)\}_{t>0}$. Let $\widehat{\mathbf{P}} = (\widetilde{P}_{ff'})_{|\mathcal{F}| \times |\mathcal{F}|}$ denote the one-step transition matrix of Ψ . By the uniformization technique, we have

$$\widehat{\mathbf{P}} = \mathbf{I} + \rho^{-1} \mathbf{Q} \tag{9}$$

where I is the identity matrix. Recall that the time required to make a transition from configuration f has an exponential

³With a slight abuse of notation, we also use f to denote the index of configuration f in the transition rate matrix.

distribution with rate given by (7). With the MAD-P/MM algorithm employed, $\zeta \sum_{n \in [N]} |Q_n^f| + (1 - \zeta) \sum_{n \in [N]} |C_n^f|$ in (7) is equal to $N \cdot R$ for any configuration f, where $R \triangleq \zeta (LW - N) + (1 - \zeta)(M - 1)$. Then, we choose

$$\rho = \alpha NR \sum_{f \in \mathcal{F}} \exp(-U(f))$$

From Algorithm 1, it can be obtained that $Q_{ff'}^{\text{cell}} = \alpha \zeta \exp(-U(f))$, where $Q_{ff'}^{\text{cell}}$ denotes the transition rate from f to some different f' due to one of the drone sBSs moving to a new cell; $Q_{ff'}^{\text{channel}} = \alpha(1-\zeta)\exp(-U(f))$, where $Q_{ff'}^{\text{channel}}$ denotes the transition rate from f to some different f' due to one of the drone sBSs hopping to a new channel. Accordingly, we have $\widehat{P}_{ff'}^{\text{cell}} = [(\alpha \zeta \exp(-U(f)))/\rho]$ and $\widehat{P}_{ff'}^{\text{channel}} = [(\alpha(1-\zeta)\exp(-U(f)))/\rho]$ for any f' that f could possibly transit to.

Next, we construct the Markov chain Ψ as follows according to (9). When the current state of Ψ is *f*:

- choose a drone sBS n* from [N] uniformly at random [i.e., with probability (1/N)];
- 2) with probability $[(\alpha NR \exp(-U(f)))/\rho]$, it is allowed to change state; otherwise, n^* remains its configuration;
- 3) when it is allowed to change state, do the following: with probability $[(\zeta (LW N))/R]$, drone n^* randomly and uniformly moves to one of the cells in $Q_{n^*}^f$; otherwise, drone n^* randomly and uniformly switches to one of the channels in $C_{n^*}^f$.

As it can be seen, by the above process, we indeed have a Markov chain Ψ according to (9). Next, we bound the mixing time of $\{\psi(t)\}_{t>0}$ through the mixing time analysis of Ψ given that $\{\psi(t)\}_{t>0}$ is represented directly as a discrete-time Markov chain Ψ with transitions governed by an independent Poisson process with rate ρ .

2) *Coupling:* We now apply the path coupling techniques [32], [37] to construct a coupling of Ψ for the mixing time analysis. A coupling of the Markov chain Ψ on the state space \mathcal{F} is a pair process $(\Psi_t, \widetilde{\Psi}_t)$ on $\mathcal{F} \times \mathcal{F}$, such that: 1) each of (Ψ_t, \cdot) and $(\cdot, \widetilde{\Psi}_t)$, viewed separately, is a faithful copy of the Markov chain Ψ and 2) if $\Psi_t = \widetilde{\Psi}_t$, we have $\Psi_{t+1} = \widetilde{\Psi}_{t+1}$. By the path coupling techniques in [37], we only need to construct a one-step coupling starting with any two adjacent states on a path, i.e., to construct $(\Psi_0, \widetilde{\Psi}_0) \rightarrow (\Psi_1, \widetilde{\Psi}_1)$, where Ψ_0 can make a transition to $\widetilde{\Psi}_0$ in one step.

Next, we define the Hamming distance H(f, f') between any two states f and f' in \mathcal{F} , which is simply the number of drone sBSs $n \in [N]$ such that $(q_n^f, c_n^f) \neq (q_n^{f'}, c_n^{f'})$.⁴ Initially, we have $H(\Psi_0, \tilde{\Psi}_0) = 1$, where Ψ_0 and $\tilde{\Psi}_0$ are two adjacent states on a path, i.e., they only differ by the cell configurations or the channel configurations of the same drone sBS. With path coupling, the following lemma shows that the Hamming distance between Ψ_1 and $\tilde{\Psi}_1$ decreases in expectation compared to $H(\Psi_0, \tilde{\Psi}_0)$, i.e., $\mathbb{E}[H(\Psi_1, \tilde{\Psi}_1)|\Psi_0, \tilde{\Psi}_0] < H(\Psi_0, \tilde{\Psi}_0)$.

Lemma 1: With path coupling under the MAD-P/MM algorithm, for any pair of Ψ_0 and $\widetilde{\Psi}_0$ (adjacent states) on $\mathcal{F} \times \mathcal{F}$,

⁴We have
$$(q_n^f, c_n^f) \neq (q_n^{f'}, c_n^{f'})$$
 if $q_n^f \neq q_n^{f'}$ or $c_n^f \neq c_n^{f'}$ or both.

we have

$$\mathbb{E}\left[H\left(\Psi_{1},\,\widetilde{\Psi}_{1}\right)|\Psi_{0},\,\widetilde{\Psi}_{0}\right] \leq (1-B)H\left(\Psi_{0},\,\widetilde{\Psi}_{0}\right)$$

where 0 < B < 1, specifically

$$B = \frac{(N-1)\left[1 + \kappa - \exp(2\beta(\Theta_{\max} - \Theta_{\min}))\right]}{N|\mathcal{F}|\exp(\beta(\Theta_{\max} - \Theta_{\min}))}$$

given that $\kappa = \min\{[(\zeta (LW - 2N + 2))/((N - 1)R)], [((1 - \zeta)M)/((N - 1)R)]\}, 0 < \beta < [(\ln(1 + \kappa))/(2(\Theta_{\max} - \Theta_{\min}))],$ and N < (LW/2) + 1.

Proof: See the Appendix for more details. *Theorem 2 (Rapid Mixing):* Under the MAD-P/MM algorithm, the mixing time of the planning process $\{\psi(t)\}_{t>0}$ is upper bounded as follows:

$$\tau_{\min}(\epsilon) \le \frac{\frac{1}{\alpha R} \exp(\beta (2\Theta_{\max} - \Theta_{\min})) \ln \frac{2N}{\epsilon}}{(N-1) \left[1 + \kappa - \exp(2\beta (\Theta_{\max} - \Theta_{\min}))\right]}$$

given that $\kappa = \min\{[(\zeta (LW - 2N + 2))/((N - 1)R)], [((1 - \zeta)M)/((N - 1)R)]\}, 0 < \beta < [(\ln(1 + \kappa))/(2(\Theta_{\max} - \Theta_{\min}))],$ and N < (LW/2) + 1.

Proof: For any $(\Psi_t, \Psi_t) \in \mathcal{F} \times \mathcal{F}$, applying Lemma 1 iteratively, we have

$$\mathbf{Pr}\{\Psi_t \neq \tilde{\Psi}_t\} = \mathbf{Pr}\{H(\Psi_t, \tilde{\Psi}_t) \ge 1\}$$

$$\leq \mathbb{E}[H(\Psi_t, \tilde{\Psi}_t)] \le (1-B)^t \cdot \operatorname{diam}(\mathcal{F})$$

where diam(\mathcal{F}) is the diameter of \mathcal{F} , i.e., the maximum of the minimum number of transitions required to go from Ψ to Ψ' over all pairs of positive-recurrent states $\Psi, \Psi' \in \mathcal{F}$. Under the MAD-P/MM algorithm, diam(\mathcal{F}) $\leq 2N$.

Therefore, for the Markov chain Ψ according to (9) and followed by the coupling lemma in [37] (Lemma 1), we have

$$d_{\mathrm{TV}}\left(\widehat{\mathbf{P}}^{t}(\Psi_{0},\cdot),\widehat{\mathbf{P}}^{t}(\widetilde{\Psi}_{0},\cdot)\right) \leq \mathbf{Pr}\left\{\Psi_{t}\neq\widetilde{\Psi}_{t}\right\} \leq 2N(1-B)^{t}$$

where $\widehat{\mathbf{P}}^{t}(\Psi_{0}, \cdot)$ ($\widehat{\mathbf{P}}^{t}(\widetilde{\Psi}_{0}, \cdot)$) denotes the *t*-step transition probability distribution starting from Ψ_{0} ($\widetilde{\Psi}_{0}$).

Consider the Markov chain $\{\psi(t)\}_{t>0}$. Recall that $\mathbf{P}_{f,\cdot}(t)$ the probability distribution over \mathcal{F} at time *t* given that the initial state is *f*, and $\mathbf{p}^* = (p_f^*)_{f \in \mathcal{F}}$ given by (1). Hence,

$$d_{\mathrm{TV}}(\mathbf{P}_{f,\cdot}(t), \boldsymbol{p}^*) = d_{\mathrm{TV}}\left(\sum_{k=0}^{\infty} \widehat{\mathbf{P}}^k(f, \cdot) \frac{\exp(-\rho t)(\rho t)^k}{k!}, \boldsymbol{p}^*\right)$$
$$\stackrel{(a)}{\leq} \sum_{k=0}^{\infty} \frac{\exp(-\rho t)(\rho t)^k}{k!} d_{\mathrm{TV}}\left(\widehat{\mathbf{P}}^k(f, \cdot), \boldsymbol{p}^*\right)$$
$$\stackrel{\leq}{\leq} 2N \sum_{k=0}^{\infty} \frac{\exp(-\rho t)(\rho(1-B)t)^k}{k!}$$
$$= 2N \exp(-\rho Bt)$$

where (a) is due to Jensen's inequality. Therefore, we have

$$\tau_{\min}(\epsilon) \leq \frac{\ln \frac{2N}{\epsilon}}{\rho B} \leq \frac{\frac{1}{\alpha R} \exp(\beta (2\Theta_{\max} - \Theta_{\min})) \ln \frac{2N}{\epsilon}}{(N-1) \left[1 + \kappa - \exp(2\beta (\Theta_{\max} - \Theta_{\min}))\right]}$$

where $\kappa = \min\{[(\zeta (LW - 2N + 2))/((N-1)R)], [((1 - \zeta)M)/((N-1)R)]\}.$

From Theorem 2, it can be seen that the mixing time decreases with the increase of α . In other words, the planning



Fig. 2. Impact of α and β on the algorithm performance. (a) MAD-P/MM, $\alpha = 1$. (b) MAD-P/LM, $\alpha = 1$. (c) MAD-P/LM, $\beta = 15$.

process $\{\psi(t)\}_{t>0}$ rapidly converges to its stationary distribution if α goes large. This is because, as discussed before, increasing α can speed up the transitions among configurations. Note that we require N < (LW/2) + 1 to have the rapid mixing of $\{\psi(t)\}_{t>0}$. Intuitively, compared to the total number of cells *LW*, if *N* is relatively large, it typically takes longer to transit between nonadjacent configurations since drones do not have much freedom to move.

C. Effects of Measurement Error

In the MAD-P algorithm, the ground-air throughput (γ_n^f) , for any drone sBS n under any configuration f) should be measured in real time, based on which the transition rates that drive the planning process can be determined [see (2)]. However, measurements of ground-air throughput could be inaccurate so that the algorithm would compute the planning based on a different set of transition rates, leading to a performance gap toward the optimal solution p^* . It is important to understand how measurement errors affect the algorithm performance and hence further characterize the performance gap. For simplicity, based on the throughput measurements from all drone sBSs under f, the broadcasted weighted system-wide normalized utility $\widehat{U}(f)$ by the central station undergoes random errors in a bounded region $[-\delta_f, \delta_f]$ compared to the true utility U(f), where δ_f is positive and dependent only on f. Hence, we have $U(f) \in [U(f) - \delta_f, U(f) + \delta_f]$. Further, we assume that $\widehat{U}(f)$ takes only $2n_f + 1$ discrete values, i.e.,

$$\begin{bmatrix} U(f) - \delta_f, U(f) - \frac{n_f - 1}{n_f} \delta_f, \dots, U(f) - \frac{1}{n_f} \delta_f, U(f) \\ U(f) + \frac{1}{n_f} \delta_f, \dots, U(f) + \frac{n_f - 1}{n_f} \delta_f, U(f) + \delta_f \end{bmatrix}$$

following certain probability distribution p_f^{error} , where n_f is a positive integer only dependent on f. With the above modeling, the performance gap in the presence of measurement errors is bounded by the following theorem.

Theorem 3: Under the MAD-P algorithm, the performance gap on the expected system utility in the presence of measurement errors is given by

$$\begin{aligned} \left| \bar{\boldsymbol{p}} \boldsymbol{U}^{T} - \boldsymbol{p}^{*} \boldsymbol{U}^{T} \right| &= \left| \sum_{f \in \mathcal{F}} \left(\bar{p}_{f} - p_{f}^{*} \right) \boldsymbol{U}(f) \right| \\ &\leq 2 \Theta_{\max} (1 - \exp(-2\beta \delta_{\max})) \end{aligned}$$

where $\bar{p} = (\bar{p}_f)_{f \in \mathcal{F}}$ denotes the stationary distribution of $\{\psi(t)\}_{t>0}$ in the present of measurement errors, $U = (U(f))_{f \in \mathcal{F}}$ denotes the vector of weighted system-wide normalized utilities, and $\delta_{\max} = \max_{f \in \mathcal{F}} \delta_f$.

The proof of Theorem 3 is omitted here since it is almost the same to the proof given by [38]. From Theorem 3, it can be seen that the performance gap decreases exponentially fast as the measurement errors diminish.

VI. SIMULATION RESULTS

In this section, we present three sets of simulations to demonstrate the performance of the proposed MAD-P algorithm, and how different design parameters (ζ , α , and β) and measurement errors affect its performance.

In the first set of simulations, we consider a target area consisting of 4×3 square cells, each of which is 100×100 m² in size. There are N = 4 drone sBSs dispatched to the target area with a height of 150 m, and there are three channels available for use, each of 15 MHz bandwidth. We consider that a fixed number of long-lived data flows are distributed in the target area, and the data traffic demand does not change over time but can be very different across cells. To compute the capacity vector of drone sBSs, a uniform ground-air transmission power of 20 dbm is used. In addition, we consider a logarithmic function as the utility function, and we set $\zeta = 0.5$.

The impact of design parameters α and β on the algorithm performance is shown in Fig. 2. The performance ratio of system utility, defined as the ratio of the running average system-wide normalized utility under the MAD-P algorithm and the optimal system-wide normalized utility obtained by the exhaustive search, is plotted with respect to the number of reconfiguration run in the simulation. As it can be seen in Fig. 2(a) and (b), with the increase of β , the performance ratio improves (i.e., the gap to the optimality is reduced) since the time fraction that the system spent in the "better" configurations increases according to (1). In Fig. 2(c), it is shown that under the MAD-P/LM algorithm the number of reconfigurations required before convergence can be large when α is small. In other words, as shown by the analytic results (Theorem 2), the convergence of the system-wide normalized utility speeds up as α goes large.

To see how ζ affects the algorithm performance, in the second set of simulations, we consider a scenario where L = 6,



Fig. 3. Impact of ζ on the algorithm performance. (a) MAD-P/MM. (b) MAD-P/LM.



Fig. 4. Impact of measurement errors on the algorithm performance. (a) MAD-P/MM. (b) MAD-P/LM.

W = 5, M = 3, N = 4, $\alpha = 1$, and $\beta = 15$, with all other parameters same to those in the first set of simulations. In this particular scenario, since the number of dispatched drone sBSs is close to the number of available channels, but much less than the number of cells in the target area, it will be more rewarding to explore the cells as much as possible so as to identify the most demanding cells very quickly. Fig. 3(a) and (b) shows that the convergence of the system-wide normalized utility speeds up by selecting a value of ζ that is close to one (i.e., more cell explorations).

We perform the third set of simulations to evaluate the impact of measurement errors on the algorithm performance, where the setting is the same to that in the first set of simulations with $\alpha = 1$ and $\beta = 15$. In addition, we introduce a uniformly generated random error when the throughput vector is measured. Fig. 4(a) and (b) shows the performance ratio of system utility in presence of measurement errors under two algorithms, respectively. As it can be seen, the measurement errors lead to an increase of utility gap. However, from the analytic results (Theorem 3), it is clear that the performance gap is upper bounded.

VII. CONCLUSION

In this paper, we considered the problem of planning a set of drone small cells operating on multiple channels in a target area to provide access to the backbone/core network. We have formulated the drone cell planning to a combinatorial network utility maximization problem, and then proposed a MAD-P algorithm to solve the problem in a distributed fashion during flight with minimal centralized control. We proved that the MAD-P algorithm is asymptotic optimal, and derived an upper bound of the mixing time that captures the speed of convergence of the dynamic planning process. We also derived an upper bound of the performance gap in presence of measurement errors. Simulation results have validated our analytical results and demonstrated the effectiveness of our algorithm.

APPENDIX

PROOF OF LEMMA 1

Denote $\Psi_0 = f^0 = (q_n^0, c_n^0)_{n \in [N]}$ and $\widetilde{\Psi}_0 = \widetilde{f}^0 = (\widetilde{q}_n^0, \widetilde{c}_n^0)_{n \in [N]}$, and we first assume Ψ_0 and $\widetilde{\Psi}_0$ (adjacent states) only differ by the cell configurations of the same drone sBS, say drone sBS 1 without loss of generality, we first assume Ψ_0 and $\widetilde{\Psi}_0$ (adjacent states) only differ by the cell configurations of drone sBS 1, i.e., $(q_n^0, c_n^0) = (\widetilde{q}_n^0, \widetilde{c}_n^0)$, $\forall n \in [2, N], q_1^0 \neq \widetilde{q}_1^0$, and $c_1^0 = \widetilde{c}_1^0$.

Next, we define the coupling at state $(\Psi_0, \widetilde{\Psi}_0)$ by choosing the next state $(\Psi_1, \widetilde{\Psi}_1)$ according to the following procedure.

Step 1: Choose drone sBS n^* from [N] uniformly at random, and both Ψ and $\widetilde{\Psi}$ update the same n^* .

Step 2: Do the following if $n^* \neq 1$, and go to step 3 otherwise.

According to the way of constructing the Markov chain Ψ , drone sBS n^* may remain its configuration in the next step or move to a new cell or hop to a new channel in the next step. We define the decision space of n^* conditional on f^0 is $\Omega(n^*|f^0) = \{\text{stay}, \mathcal{Q}_{n^*}^{f^0}, \mathcal{C}_{n^*}^{f^0}\}$, and the probability distribution $p_{n^*} = (p_{n^*}(\omega))_{\omega \in \Omega(n^*|f^0)}$, where

$$p_{n^*}(\omega) = \begin{cases} 1 - \frac{\alpha NR \exp(-U(f^0))}{\rho}, & \omega = \text{stay} \\ \frac{\alpha \zeta N \exp(-U(f^0))}{\rho}, & \omega \in \mathcal{Q}_{n^*}^{f^0} \\ \frac{\alpha (1-\zeta)N \exp(-U(f^0))}{\rho}, & \omega \in \mathcal{C}_{n^*}^{f^0}. \end{cases}$$

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Similarly, we define the decision space of n^* conditional on \tilde{f}^0 is $\Omega(n^*|\tilde{f}^0) = \{\text{stay}, \mathcal{Q}_{n^*}^{\tilde{f}^0}, \mathcal{C}_{n^*}^{\tilde{f}^0}\}$, and the probability distribution $\tilde{p}_{n^*} = (\tilde{p}_{n^*}(\omega))_{\omega \in \Omega(n^*|\tilde{f}^0)}$, where

$$\widetilde{p}_{n^*}(\omega) = \begin{cases} 1 - \frac{\alpha N R \exp(-U(\widetilde{f}^0))}{\rho}, & \omega = \operatorname{stay} \\ \frac{\alpha \zeta N \exp(-U(\widetilde{f}^0))}{\rho}, & \omega \in \mathcal{Q}_{n^*}^{\widetilde{f}^0} \\ \frac{\alpha (1-\zeta) N \exp(-U(\widetilde{f}^0))}{\rho}, & \omega \in \mathcal{C}_{n^*}^{\widetilde{f}^0}. \end{cases}$$

Based on p_{n^*} over $\Omega(n^*|f^0)$ and \widetilde{p}_{n^*} over $\Omega(n^*|\widetilde{f}^0)$, we define p'_{n^*} and \widetilde{p}'_{n^*} over the same sample space $\Omega(n^*) = \Omega(n^*|f^0) \cup \Omega(n^*|\widetilde{f}^0) = \{\text{stay}, \mathcal{Q}_{n^*}^{f^0} \cup \mathcal{Q}_{n^*}^{\widetilde{f}^0}, \mathcal{C}_{n^*}^{f^0}\}$. Note that $\mathcal{C}_{n^*}^{f^0} = [M] \setminus \mathcal{C}_{n^*}^{0} = [M] \setminus \widetilde{\mathcal{C}}_{n^*}^{0}$ since $\mathcal{C}_{n^*}^{0} = \widetilde{\mathcal{C}}_{n^*}^{0}$. Specifically, we have

$$p_{n^*}'(\omega) = \begin{cases} p_{n^*}(\omega) & \forall \omega \in \Omega(n^*|f^0) \\ 0 & \forall \omega \in \Omega(n^*) \setminus \Omega(n^*|f^0) \end{cases}$$
(10)

$$\widetilde{p}_{n^*}'(\omega) = \begin{cases} \widetilde{p}_{n^*}(\omega) & \forall \omega \in \Omega(n^*|f^0) \\ 0 & \forall \omega \in \Omega(n^*) \setminus \Omega(n^*|\widetilde{f}^0). \end{cases}$$
(11)

We further define the following three probability distributions over the sample space $\Omega(n^*)$:

$$\boldsymbol{p}_{n^*}^{\min} = \frac{\left(\min\{p'_{n^*}(\omega), \widetilde{p}'_{n^*}(\omega)\}\right)_{\omega \in \Omega(n^*)}}{1 - d_{\mathrm{TV}}(\boldsymbol{p}'_{n^*}, \widetilde{\boldsymbol{p}}'_{n^*})}$$
(12)

$$\boldsymbol{p}_{n^{*}}^{+} = \frac{\left(\max\left\{0, p_{n^{*}}'(\omega) - \widetilde{p}_{n^{*}}'(\omega)\right\}\right)_{\omega \in \Omega(n^{*})}}{d_{\mathrm{TV}}(\boldsymbol{p}_{n^{*}}', \widetilde{\boldsymbol{p}}_{n^{*}}')}$$
(13)

$$\widetilde{p}_{n^*}^+ = \frac{\left(\max\left\{0, \widetilde{p}_{n^*}'(\omega) - p_{n^*}'(\omega)\right\}\right)_{\omega \in \Omega(n^*)}}{d_{\mathrm{TV}}(p_{n^*}', \widetilde{p}_{n^*}')}.$$
 (14)

Recall that $d_{\text{TV}}(\cdot, \cdot)$ denotes the total variation distance of two probability distributions.

Now, we are ready to update drone sBSs n^* in both Markov chains Ψ and $\widetilde{\Psi}$ for $(\Psi_1, \widetilde{\Psi}_1)$.

- 1) With probability $1 d_{\text{TV}}(\mathbf{p}'_{n^*}, \widetilde{\mathbf{p}}'_{n^*})$, select a configuration ω from $\Omega(n^*)$ according to $\mathbf{p}_{n^*}^{\min}$ given in (12), and both Ψ and $\widetilde{\Psi}$ update n^* to the new configuration ω .
- 2) Otherwise, Ψ and $\tilde{\Psi}$ update n^* independently. Specifically, Ψ updates n^* according to $p_{n^*}^+$ given in (13), and $\tilde{\Psi}$ updates n^* according to $\tilde{p}_{n^*}^+$ given in (14).

and $\widetilde{\Psi}$ updates n^* according to $\widetilde{p}_{n^*}^+$ given in (12), step 3: Given that $n^* = 1$, $q_1^0 \neq \widetilde{q}_1^0$, and $c_1^0 = \widetilde{c}_1^0$, we define two probability distributions p'_1 and \widetilde{p}'_1 over the sample space $\Omega(1) = \{(q_1^0, c_1^0), (\widetilde{q}_1^0, c_1^0), \mathcal{Q}_1^{f^0} \cap \mathcal{Q}_1^{\widetilde{f}^0}\}$. Specifically,

$$p_{1}'(\omega) = \begin{cases} 1 - \frac{\alpha NR \exp(-U(f^{0}))}{\rho}, & \omega = (q_{1}^{0}, c_{1}^{0}) \\ \frac{\alpha \zeta N \exp(-U(f^{0}))}{\rho}, & \omega = (\widetilde{q}_{1}^{0}, c_{1}^{0}) \\ \frac{\alpha \zeta N \exp(-U(f^{0}))}{\rho}, & \omega \in \mathcal{Q}_{1}^{f^{0}} \cap \mathcal{Q}_{1}^{\widetilde{f}^{0}} \\ \frac{\alpha (1-\zeta) N \exp(-U(f^{0}))}{\rho}, & \omega \in \mathcal{C}_{1}^{f^{0}} \\ 0, & \text{otherwise} \end{cases}$$

$$\widetilde{p}_{1}'(\omega) = \begin{cases} 1 - \frac{\alpha NR \exp(-U(\widetilde{f}^{0}))}{\rho}, & \omega = (\widetilde{q}_{1}^{0}, c_{1}^{0}) \\ \frac{\alpha \zeta N \exp(-U(\widetilde{f}^{0}))}{\rho}, & \omega = (q_{1}^{0}, c_{1}^{0}) \\ \frac{\alpha \zeta N \exp(-U(\widetilde{f}^{0}))}{\rho}, & \omega \in \mathcal{Q}_{1}^{f^{0}} \cap \mathcal{Q}_{1}^{\widetilde{f}^{0}} \\ \frac{\alpha (1-\zeta) N \exp(-U(\widetilde{f}^{0}))}{\rho}, & \omega \in \mathcal{C}_{1}^{f^{0}} \\ 0, & \text{otherwise.} \end{cases}$$

Similarly, we further define the following three probability distributions over the sample space $\Omega(1)$:

$$\boldsymbol{p}_{1}^{\min} = \frac{\left(\min\{p_{1}'(\omega), \widetilde{p}_{1}'(\omega)\}\right)_{\omega \in \Omega(1)}}{1 - d_{\mathrm{TV}}(\boldsymbol{p}_{1}', \widetilde{\boldsymbol{p}}_{1}')}$$
(15)

$$\boldsymbol{p}_{1}^{+} = \frac{\left(\max\{0, p_{1}^{\prime}(\omega) - \widetilde{p}_{1}^{\prime}(\omega)\}\right)_{\omega \in \Omega(1)}}{d_{\mathrm{TV}}(\boldsymbol{p}_{1}^{\prime}, \widetilde{\boldsymbol{p}}_{1}^{\prime})}$$
(16)

$$\widetilde{p}_{1}^{+} = \frac{\left(\max\left\{0, \widetilde{p}_{1}^{\prime}(\omega) - p_{1}^{\prime}(\omega)\right\}\right)_{\omega \in \Omega(1)}}{d_{\mathrm{TV}}\left(p_{1}^{\prime}, \widetilde{p}_{1}^{\prime}\right)}.$$
(17)

Now, we are ready to update drone sBSs n^* in both Markov chains Ψ and $\widetilde{\Psi}$ for $(\Psi_1, \widetilde{\Psi}_1)$.

- 1) With probability $1 d_{\text{TV}}(\mathbf{p}'_1, \widetilde{\mathbf{p}}'_1)$, select a configuration ω from $\Omega(1)$ according to \mathbf{p}_1^{\min} given in (15), and both Ψ and $\widetilde{\Psi}$ update drone sBS 1 to the new configuration ω .
- 2) Otherwise, Ψ and $\tilde{\Psi}$ update drone sBS 1 independently. Specifically, Ψ updates drone sBS 1 according to p_1^+ given in (16), and $\tilde{\Psi}$ updates drone sBS 1 according to \tilde{p}_1^+ given in (17).

It is not hard to show that the above procedure leads to a valid coupling. Initially, we have $H(\Psi_0, \tilde{\Psi}_0) = 1$. With path coupling, we next show that the Hamming distance between Ψ_1 and $\tilde{\Psi}_1$ decreases in expectation compared to $H(\Psi_0, \tilde{\Psi}_0)$, i.e., $\mathbb{E}[H(\Psi_1, \tilde{\Psi}_1)|\Psi_0, \tilde{\Psi}_0] < H(\Psi_0, \tilde{\Psi}_0)$.

By the above coupling procedure, it can be seen that when $n^* = 1$, we have $\Psi_1 = \widetilde{\Psi}_1$ if drone sBS 1 remains its configuration from Ψ_0 to Ψ_1 and moves from \widetilde{q}_1^0 ($\widetilde{\Psi}_0$) to q_1^0 ($\widetilde{\Psi}_1$)); or drone sBS 1 remains its configuration from $\widetilde{\Psi}_0$ to $\widetilde{\Psi}_1$ and moves from q_1^0 (Ψ_0) to \widetilde{q}_1^0 (Ψ_1); or both chains update drone sBS 1 to a new cell in $\mathcal{Q}_1^{f^0} \cap \mathcal{Q}_1^{\widetilde{f}^0}$. Note that $|\mathcal{Q}_1^{f^0} \cap \mathcal{Q}_1^{\widetilde{f}^0}| = LW - N - 1$. Hence, we have

$$\mathbb{E}\left[H\left(\Psi_{1},\widetilde{\Psi}_{1}\right)-1|\Psi_{0},\widetilde{\Psi}_{0},n^{*}=1\right]$$

$$=-\sum_{\omega\in\Omega(1)\setminus\mathcal{C}_{1}^{f^{0}}}\left(1-d_{\mathrm{TV}}\left(p_{1}^{\prime},\widetilde{p}_{1}^{\prime}\right)\right)p_{1}^{\min}(\omega)$$

$$=-\sum_{\omega\in\Omega(1)\setminus\mathcal{C}_{1}^{f^{0}}}\min\{p_{1}^{\prime}(\omega),\widetilde{p}_{1}^{\prime}(\omega)\}$$

$$\leq-\frac{\zeta(LW-N+1)\exp(-\beta(\Theta_{\max}-\Theta_{\min}))}{R|\mathcal{F}|}.$$
 (18)

When $n^* \in [2, N]$, with probability $1 - d_{\text{TV}}(\mathbf{p}'_{n^*}, \widetilde{\mathbf{p}}'_{n^*})$, $H(\Psi_1, \widetilde{\Psi}_1) = 1$; otherwise, $H(\Psi_1, \widetilde{\Psi}_1) = 2$. Then, we have $\mathbb{E}[H(\Psi_1, \widetilde{\Psi}_1) - 1 | \Psi_0, \widetilde{\Psi}_0, n^* \in [2, N]]$ $= d_{\text{TV}}(\mathbf{p}'_{n^*}, \widetilde{\mathbf{p}}'_{n^*}) = 1 - \sum_{\omega \in \Omega(n^*)} \min\{p'_{n^*}(\omega), \widetilde{p}'_{n^*}(\omega)\}.$ (19)

From the MAD-P/MM algorithm, it can be obtained that for any $n^* \in [2, N], |\mathcal{Q}_{n^*}^{f^0}| = |\mathcal{Q}_{n^*}^{f^0}| = LW - N$, and $|\mathcal{Q}_{n^*}^{f^0} \cap \mathcal{Q}_{n^*}^{f^0}| = LW - N - 1$. Therefore,

$$\sum_{\omega \in \Omega(n^*)} \min \left\{ p'_{n^*}(\omega), \widetilde{p}'_{n^*}(\omega) \right\}$$

$$\geq 1 - \frac{\exp(\beta(\Theta_{\max} - \Theta_{\min}))}{|\mathcal{F}|}$$

$$+ \frac{\zeta(LW - N - 1)\exp(-\beta(\Theta_{\max} - \Theta_{\min}))}{R|\mathcal{F}|}$$

$$+ \frac{(1 - \zeta)(M - 1)\exp(-\beta(\Theta_{\max} - \Theta_{\min}))}{R|\mathcal{F}|}$$

$$= 1 - \frac{\exp(-\beta(\Theta_{\max} - \Theta_{\min}))}{|\mathcal{F}|}$$

$$\times \left(\exp(2\beta(\Theta_{\max} - \Theta_{\min})) - \frac{R - \zeta}{R} \right). \quad (20)$$

Here, we assume that

$$\exp(\beta(\Theta_{\max} - \Theta_{\min})) < |\mathcal{F}| \tag{21}$$

since $1 - [(\exp(\beta(\Theta_{\max} - \Theta_{\min})))/(|\mathcal{F}|)]$ has to be a valid probability. However, we will immediately have (21) after we bound β shortly. By plugging (18)–(20), we bound the conditional expected Hamming distance between Ψ_1 and $\widetilde{\Psi}_1$

$$\mathbb{E}\left[H\left(\Psi_{1}, \widetilde{\Psi}_{1}\right) - 1|\Psi_{0}, \widetilde{\Psi}_{0}\right] \\
= \sum_{n \in [N]} \frac{1}{N} \mathbb{E}\left[H\left(\Psi_{1}, \widetilde{\Psi}_{1}\right) - 1|\Psi_{0}, \widetilde{\Psi}_{0}, n^{*} = n\right] \\
\leq \frac{1}{N} \cdot \frac{1}{|\mathcal{F}|} (N - 1) \exp(-\beta(\Theta_{\max} - \Theta_{\min})) \\
\times \left[\exp(2\beta(\Theta_{\max} - \Theta_{\min})) - \left(1 + \frac{\zeta(LW - 2N + 2)}{(N - 1)R}\right)\right].$$
(22)

$$B_1 \triangleq \frac{(N-1)\left[1 + \frac{\zeta(LW - 2N + 2)}{(N-1)R} - \exp(2\beta(\Theta_{\max} - \Theta_{\min}))\right]}{N|\mathcal{F}|\exp(\beta(\Theta_{\max} - \Theta_{\min}))}$$

which is positive if the following two conditions hold:

$$N < \frac{LW}{2} + 1$$

$$0 < \beta < \frac{1}{2(\Theta_{\max} - \Theta_{\min})} \ln\left(1 + \frac{\zeta (LW - 2N + 2)}{(N - 1)R}\right).$$
(24)

Due to (24), we certainly have (21). Therefore, for any pair of Ψ_0 and $\widetilde{\Psi}_0$ (adjacent states) only differing by the cell configurations of the same drone sBS, it follows that:

$$\mathbb{E}\left[H\left(\Psi_{1},\widetilde{\Psi}_{1}\right)|\Psi_{0},\widetilde{\Psi}_{0}\right] \leq 1 - B_{1} = (1 - B_{1})H\left(\Psi_{0},\widetilde{\Psi}_{0}\right)$$

where $0 < B_1 < 1$.

Follow the above analysis, similarly, for any pair of Ψ_0 and $\widetilde{\Psi}_0$ (adjacent states) only differing by the channel configurations of the same drone sBS, we have

$$\mathbb{E}\left[H\left(\Psi_{1}, \widetilde{\Psi}_{1}\right)|\Psi_{0}, \widetilde{\Psi}_{0}\right] \leq 1 - B_{2} = (1 - B_{2})H\left(\Psi_{0}, \widetilde{\Psi}_{0}\right)$$

where

$$B_2 \triangleq \frac{(N-1) \left[1 + \frac{(1-\zeta)M}{(N-1)R} - \exp(2\beta(\Theta_{\max} - \Theta_{\min})) \right]}{N|\mathcal{F}|\exp(\beta(\Theta_{\max} - \Theta_{\min}))}$$

which is positive when

$$0 < \beta < \frac{1}{2(\Theta_{\max} - \Theta_{\min})} \ln\left(1 + \frac{(1 - \zeta)M}{(N - 1)R}\right).$$
(25)

Now, we choose any β which satisfies

$$0 < \beta < \frac{\ln(1+\kappa)}{2(\Theta_{\max} - \Theta_{\min})}$$
(26)

where $\kappa \triangleq \min\{[(\zeta (LW - 2N + 2))/((N - 1)R)], [((1 - \zeta)M)/((N - 1)R)]\}$. We further define $B \triangleq \min\{B_1, B_2\}$. From the path coupling theorem in [37], for any $(\Psi_0, \widetilde{\Psi}_0) \in \mathcal{F} \times \mathcal{F}$, we have

$$\mathbb{E}\left[H\left(\Psi_{1}, \widetilde{\Psi}_{1}\right)|\Psi_{0}, \widetilde{\Psi}_{0}\right] \leq (1 - B)H\left(\Psi_{0}, \widetilde{\Psi}_{0}\right)$$
(27)

where 0 < B < 1. Thus, Lemma 1 follows.

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